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Institution and Development Revisited: A Nonparametric Approach

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Institution and Development Revisited: A Nonparametric Approach

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Abstract

The paper uses nonparametric methodology to examine the role of institutions in understanding differential levels of development across countries. This technique estimates first order derivatives for every country allowing a deeper look into the impact of institutions on development. The preliminary cross-country findings show that (i) institutional quality positively and significantly increases development quality; and (ii) results remain 'robust' for different model specifications and choice of additional control variables. The analysis is carried out for a set of 102 countries over 1980 to 2004. Similar to parametric results established in the literature, the nonparametric analysis lends further support to the view that institutions matter in the context of economic policies and geographic factors.

Keywords: Development, Institutions, Geography, Openness, Principal component, Nonparametric analysis

JEL Classification Numbers: C3, O10, O57, P51, R11

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1. Introduction

Do institutions cause differential levels of development across countries? Should the development agenda of an under-developed country be directed towards building institutions with standards similar to those of developed countries?¹ What effects do institutions have on indicators of development? Answers to these questions are relevant for policy makers and planners worldwide.

The relevant literature states that geography, economic policy and institutions are the three most important determinants of a country's economic performance. The *endowment hypothesis* states that geography/ biogeographic or climatic conditions explain cross-country differences in economic performances (Diamond (1997), Gallup et. al. (1998), Masters and McMillan (2001) and Hibbs and Olsson (2005, 2004)).² This body of literature suggests that, "Institutions Don't Rule" (Sachs (2003)). According to the *policy hypothesis*, efficient resource allocation by economic policy is responsible for faster economic growth (Sachs and Warner (1995), Edwards (1998), Frankel and Romar (1999), Dollar and Kraay (2001, 2003) and Wacziarg and Welch (2003)).³ The *institutions hypothesis* advocates that quality of institutions trumps both geography as well as policy in determining a country's level of development (Acemoglu, Johnson and Robinson (2001), Rodrik, Subramanian and Trebbi (2004), Easterly and Levine (2003) and Basu (2008)).

The purpose of our paper is to further investigate the *institutions hypothesis*. We use two innovative measures of development and institutional quality. Our paper explores the complexities of development and its interactions with institutions, geography and economic

¹ There is no established convention for the designation of "developed" and "developing" countries or areas in the United Nations system. In common practice, Japan in Asia, Canada and the United States in northern America, Australia and New Zealand in Oceania, and Europe are considered "developed" regions or areas. For details refer to the United Nations Statistics Division. Table A1 gives a complete list and classification of countries used in the paper.

² Gallup et. al control for macroeconomic policies, while Hibbs and Olsson (2005, 2004), control for institutions and economic policies. They find that only geography matters for economic performance.

³ However, Stiglitz (1999), Rodriguez and Rodrik (2000) and Muqtada (2003) question the effectiveness of trade reform and macroeconomic polices on the economy, in the absence of institutional support.

policies within the framework of nonparametric and semiparametric analysis. We focus on estimating the relationship between institutional and development quality in a data driven specification free manner.

The existing body of literature uses single indicators such as, GDP per capita as a proxy for development or the rule of law and property rights to measure institutional quality. For our analysis, we use two indices, the development quality index (DQI) and the institutional quality index (IQI), from the methodology proposed in Nagar and Basu (2002). These indices are capable of capturing a broader range of issues related to development and institutions. According to Acemoglu, Johnson and Robinson (2001), institutions positively influence GDP per capita, by securing property rights.⁴ Their estimates suggest that, geography does not cause variations in GDP per capita. Rodrik, Subramanian and Trebbi (2004) argue that institutions dominate geography and trade policies in influencing income levels around the world. Easterly and Levine (2003) show that economic policy has no effect on GDP per capita after controlling for institutions and geography effects a country's GDP indirectly via institutions. According to them, only institutions can explain cross-country variations in GDP per capita and institutional indicators would be a sufficient statistic to account for variations in economic performance.⁵ These studies argue that "Institutions Rule" over geography and economic policy. Basu (2008) strongly supports the importance of institutions in the context of specific economic policy mixes and geography. These highly quoted studies use parametric estimation techniques. The nature of the relationship between institutions and development is the heart of the issue. Therefore, we employ parametric and semiparametric estimators to investigate the relationship between various institutional and development indicators, in a cross-country context.

⁴ In Acemoglu, Johnson and Robinson (2001), property rights are measured as average protection against expropriation risk.

⁵ Bardhan (2005) argues that institutions could play an important role in determining economic performance, but question still remains "Institutions matter, but which ones?"

We estimate a nonparametric and semiparametric model to capture the relationship between institutional quality and development quality for 102 countries from 1980 to 2004. This flexible estimation strategy helps us examine the development-institution relationship in a data driven manner without superimposing any *a priori* functional form restriction. We also feel that the nonparametric estimation technique can be exploited to provide a deeper understanding of the relationship of interest. The nonparametric methodology gives us an estimate of the first order derivative of the development quality with respect to the institutional quality for every country.

We now plot a course for the rest of the paper. Section 2 presents the data and estimation technique used in the paper. Section 2.1 outlines the nonparametric methodology and section 2.2 discusses the semiparametric methodology. We discuss our main findings and results in Section 3. Section 4 concludes the paper.

2. Data and Estimation

Our paper is based on 102 countries, of which 76 are developing countries, 22 are OECD countries, and 29 are least developed and small-medium size countries, as defined by United Nations and WTO respectively (See Table A1 for a complete list). We look at data of indicators from several international sources, research institutions and think-tanks (See Table A2 for data sources of the indicators used in the paper). For our analysis, we compute two indices, the development quality index (*DQI*) and the institutional quality index (*IQI*), for 102 countries and for five time intervals: 1980-1984, 1985-1989, 1990-1994, 1995-1999, and 2000-2004. We take an average of the indices from each time interval to get the index values for a country over the entire period for our analysis.

The *DQI* is calculated from three aspects of development: economic, health and knowledge. Economic development indicators are, GDP per capita in PPP international 2000

\$, telephone lines, television sets, radios, electric power consumption per capita, and energy use per capita; health development indicators are, life expectancy at birth, infant mortality rate, physicians, immunization of children, and CO₂ emissions per capita; and knowledge development indicators are, adult literacy rate, primary school enrolment rate, secondary school enrolment rate and total number of years in schools. The *DQI* is a composite index, which covers 15 indicators of development. Likewise, the *IQI* is constructed to evaluate the quality of institutions. It is also calculated from three aspects of institutional quality: economic, social and political. Economic institutional quality is a combination of legal and property rights, bureaucratic quality, corruption, democratic accountability, government stability, law and order, independent judiciary, and regulation; social institutional quality is based on press freedom, civil liberties, physical integrity index, empowerment right index, freedom of association, women's political rights, women's economic right, and women's social right; and political institutional quality depends on executive constraint, index of democracy, political rights, polity score, lower legislative, upper legislative and independent sub-federal units. The *IQI* is based on 23 indicators of quality of institutions.⁶

2.1 The Nonparametric Model

To make our analysis robust, we use four indices of institutional quality: the political institutional quality index (*PIQI*), the social institutional quality index (*SIQI*), the economic institutional quality index (*EIQI*) and the total institutional quality index (*IQI*). The geography indicator (*GEOG*) is the absolute distance of a country from the equator and the openness/ world integration indicator (*EPOL*) is a trade/GDP ratio. Most studies concerned with the relationship between development and institutions, estimate a parametric equation of the form: $DQI_i = \beta_0 + \beta_1 EPOL_i + \beta_2 GEOG_i + \beta_3 IQI_i + v_i$.

⁶ See A2 for definition and sources of DQI and IQI indicators.

To capture the relationship between institutional quality and development quality, we replace a typical parametric model with the corresponding nonparametric model. This flexible estimation strategy helps us avoid any functional form misspecification bias and enables us to explore the shape of the underlying relationship without superimposing any *a priori* functional form restriction. As a result, the development- institution relationship is examined in a data driven manner. We also feel that the nonparametric estimation technique can be exploited to provide a deeper understanding of the relationship of interest. The nonparametric methodology gives us an estimate of the slope coefficient for every country. Hence we can examine the relationship between *DQI* and the institutional quality indices: *IQI*, *PQI*, *SIQI* and *EIQI*, for each country. Equations (1) through (4) capture the nonparametric relationships between development and institutional quality indices. Here, $m_1(\cdot)$, $m_2(\cdot)$, $m_3(\cdot)$ and $m_4(\cdot)$ are unknown smooth functions; and the classical error terms are, v_{1i} , v_{2i} , v_{3i} and v_{4i} . We represent the set of equations in (1) through (4) with a general form in (5). Y_i is the development quality index for the i^{th} country, $\{X_i\}$ is the 1×3 matrix of independent variables, v_i is a stochastic disturbance term and the sample size $n = 102$.⁷

$$DQI_i = m_1(EPOL_i, GEOG_i, IQI_i) + v_{1i} \quad (1)$$

$$DQI_i = m_2(EPOL_i, GEOG_i, PIQI_i) + v_{2i} \quad (2)$$

$$DQI_i = m_3(EPOL_i, GEOG_i, SIQI_i) + v_{3i} \quad (3)$$

$$DQI_i = m_4(EPOL_i, GEOG_i, EIQI_i) + v_{4i} \quad (4)$$

$$Y_i = m(X_i) + v_i \quad (i = 1, 2, \dots, n) \quad (5)$$

We assume $m(\cdot)$ is a smooth function and the sample realizations $\{Y_i, X_i\}$ are i.i.d. Then, we can estimate $m(\cdot)$ nonparametrically using kernel methods. Details of this estimation strategy are available in Silverman (1998), Pagan and Ullah (1999) and Li and

⁷ $X_i = [EPOL_i \text{ } GEOG_i \text{ } IQI_i]$ in (1), $X_i = [EPOL_i \text{ } GEOG_i \text{ } PIQI_i]$ in (2), $X_i = [EPOL_i \text{ } GEOG_i \text{ } SIQI_i]$ in (3) and $X_i = [EPOL_i \text{ } GEOG_i \text{ } EIQI_i]$ in (4); $v_i = v_{1i}$ in (1), $v_i = v_{2i}$ in (2), $v_i = v_{3i}$ in (3) and $v_i = v_{4i}$ in (4).

Raccine (2006).⁸ Taking a linear Taylor series expansion and considering terms up to the first order, we can re-write the local linear version of (5) as,

$$Y_i = m(x) + (X_i - x)' \beta(x) + \varepsilon_i \quad (6)$$

In (10), ε_i is the new error term that captures the remaining terms of the Taylor series expansion.⁹ In our notation, $m(\cdot)$ is the unknown conditional mean of Y_i , $\beta(\cdot)$ is a vector of the partial derivative of Y_i with respect to $\{X_i\}$, $\{X_i'\}$ is a (1×3) matrix of explanatory variables and x' is a (1×3) matrix of actual data. Re-arranging terms in (6), we get, $Y_i = Z_i \delta(x) + \varepsilon_i$, where $Z_i = [1 \ X_i']$ and $\delta(x) = [m(x) - x\beta(x) \ \beta(x)]'$. In matrix notation, $Y = Z\delta + \varepsilon$, where Y is a $(n \times 1)$ matrix of dependent variables, Z is a $(n \times 4)$ matrix of ones and independent variables, δ is a (4×1) matrix of parameters and ε is a $(n \times 1)$ matrix of stochastic errors. We minimize the weighted sum of squares of residuals, $S = (K^{1/2}Y - K^{1/2}Z\delta)'(K^{1/2}Y - K^{1/2}Z\delta)$, to obtain the local linear least squares estimates (LLLS) $\delta_{LLLS} = (Z'KZ)^{-1}(Z'KY)$, where $V(\delta_{LLLS}) = \sigma^2(Z'KZ)^{-1}(Z'KK'Z)(Z'KZ)^{-1}$. The $n \times n$ matrix of weights $K(x) = \text{Diag}(K((X_i - x)/h))$. It is well known in the literature that the choice of kernels does not influence significantly the efficiency of LLLS estimates (Silverman, 1998, Table 3.1, pg 43). The choice of window width is however crucial. Small values of h cause over smoothing and high values lead to under smoothing of the estimates. The optimum h is the one that minimizes the integrated mean squared error of $m(x)$.¹⁰ To make the LLLS estimate feasible, we estimate $\hat{\sigma}^2 = 1/(n-k-1) \sum e_i^2$, where e_i is the least squares residual from (5).

We obtain the local linear generalized least squares (LLGLS) estimates by minimizing the weighted sum of squares of residuals $S' = (Y - Z\delta)' K^{1/2} \Omega^{-1} K^{1/2} (Y - Z\delta)$ with

⁸ For a simple exposition refer to Das (2008).

⁹ We assume that $E(\varepsilon_i) = 0$ and $V(\varepsilon_i) = \sigma^2$.

¹⁰ We choose the multivariate Gaussian kernel density function $K(x) = \prod_{q=1}^3 (1/\sqrt{2\pi}) \exp(-(x_{iq} - x_q)/2h_q)^2$, $h_q = 1.06n^{-1/5} s_{xq}$ and s_{xq} is the standard error of the q^{th} independent variable.

respect to $\delta(x)$. Essentially, we apply the least squares estimation on the transformed observations $\Omega^{-1/2}K^{1/2}Y$ and $\Omega^{-1/2}K^{1/2}X$ and obtain $\delta_{LLGLS}=(Z'K^{1/2}\Omega^{-1}K^{1/2}Z)^{-1}(Z'K^{1/2}\Omega^{-1}K^{1/2}Y)$ with $V(\delta_{LLGLS})=(Z'\theta Z)^{-1}Z'\theta\Omega^{-1}\theta'Z(Z'\theta Z)^{-1}$, where $\theta=K^{1/2}\Omega^{-1}K^{1/2}$. Following the suggestions of Hinkley (1977), to make the LLGLS estimates feasible, we estimate $\hat{\Omega}=(n/(n-k))\zeta$, where a typical element of $\zeta=\text{diag}(e_1^2, e_2^2 \dots e_n^2)$.¹¹ Both LLS and LLGLS estimates are consistent. Their asymptotic properties are established in Pagan and Ullah (1999) and Li and Raccine (2006).

2.2 The Semiparametric Model

To check for robustness of our nonparametric estimates, we introduce several geographical variables into our original model. These geographical variables are, the geographical dummy for Africa (*reg_africa*), Asia (*reg_asia*) and Latin America (*reg_lac*). We consider the semiparametric model of Yatchew (2003). Equations (7) through (10) are partially linear models where the regression equation has both parametric as well as unspecified components. Here, $m_5(\cdot)$, $m_6(\cdot)$, $m_7(\cdot)$ and $m_8(\cdot)$ are unknown smooth functions and the classical error terms are, v_{5t} , v_{6t} , v_{7t} and v_{8t} . We represent the set of equations in (7) to (10) with a general form in (11). Y_i is the development quality index for the i^{th} country, $Z_i=[GEOG_i \text{ reg_africa}_i \text{ reg_asia}_i \text{ reg_lac}_i]$ is a (1×4) matrix of geography dummies, X_i is a (1×2) matrix of independent variables and v_i is a stochastic disturbance term.¹² Equation (12) is the matrix form representation of equation (11). $Y=[Y_i]$ is a (n×1) vector of dependent variables, $Z=[Z_i]$ is a (n×4) vector of geography dummies, which have a parametric relationship with Y

¹¹ It is well known that White corrected t values tend to be too large. Hence we use Hinkley's estimates for the variance covariance matrix. We also follow Horn and Duncan's (1975) suggestion of using $S = \text{diag}(s_1^2, \dots, s_n^2)$, where $s_i^2 = e_i^2/(1-m_{ii})$ and m_{ii} = diagonal element of $X(XX')^{-1}X'$. Results from Horn and Duncan's methodology are very similar to the results from Henkley (1977). So only Henley estimates are used to make the LLGLS estimates feasible.

¹² $X_i = [EPOL_i \text{ IQI}_i]$ in (7), $X_i = [EPOL_i \text{ PIQI}_i]$ in (8), $X_i = [EPOL_i \text{ SIQI}_i]$ in (9) and $X_i = [EPOL_i \text{ EIQI}_i]$ in (10); $v_i = v_{5i}$ in (7), $v_i = v_{6i}$ in (8), $v_i = v_{7i}$ in (9) and $v_i = v_{8i}$ in (10).

and $X = [X_i]$ is a $(n \times 2)$ matrix of independent variables, which have a nonparametric relationship with Y . We assume that $E(\varepsilon|Z, X)=0$ and $V(\varepsilon|Z, X)=\sigma_\varepsilon^2$.

$$DQI_i = m_5(EPOL_i, IQI_i) + \beta_0 GEOG_i + \beta_1 reg_asia + \beta_2 reg_africa + \beta_3 reg_lac + v_{5i} \quad (7)$$

$$DQI_i = m_6(EPOL_i, PIQI_i) + \beta_0 GEOG_i + \beta_1 reg_asia + \beta_2 reg_africa + \beta_3 reg_lac + v_{6i} \quad (8)$$

$$DQI_i = m_7(EPOL_i, SIQI_i) + \beta_0 GEOG_i + \beta_1 reg_asia + \beta_2 reg_africa + \beta_3 reg_lac + v_{7i} \quad (9)$$

$$DQI_i = m_8(EPOL_i, EIQI_i) + \beta_0 GEOG_i + \beta_1 reg_asia + \beta_2 reg_africa + \beta_3 reg_lac + v_{8i} \quad (10)$$

$$Y_i = m(X_i) + Z_i' \beta + v_i \quad (11)$$

$$Y = m(X) + Z\beta + \varepsilon \quad (12)$$

We suppose that the conditional mean of Z , $E(Z|X)$ is a smooth bounded function of X , say $g(X)$ and $V(Z|X) = \sigma_u^2$. Then, we may write $Z = g(X) + u$. Differencing yields, $(Y_i - Y_{i-1}) = (Z_i - Z_{i-1})\beta + (f(X_i) - f(X_{i-1})) + \varepsilon_i - \varepsilon_{i-1}$. Applying OLS to the differenced data, we get $b_{DIFF} = \Sigma(Y_i - Y_{i-1})(Z_i - Z_{i-1}) / \Sigma(Z_i - Z_{i-1})^2$. The semiparametric estimates are obtained by applying local linear least squares to the transformed model: $Y - Zb_{DIFF} = m(X) + \varepsilon$. Yatchew (2003) works out the consistency properties of the semiparametric estimates.

3: Results

For the parametric results of a similar model, we urge the reader to refer to Basu (2008). In this paper, we focus on the nonparametric model and results. We estimate the nonparametric regression functions in (1) through (4) with local linear least squares methodology outlined in section 2.1. A challenge with using nonparametric technology is that we obtain 102 estimates of first order derivatives for each independent variable. To keep the analysis manageable, we focus on the main relationship of interest, the development-institution relationship. The nonparametric estimates of first order derivatives are represented by $\delta = \{\delta_1 \delta_2 \delta_3 \delta_4\}$, where $\delta_1 = \partial DQI / \partial IQI$, $\delta_2 = \partial DQI / \partial PIQI$, $\delta_3 = \partial DQI / \partial SIQI$ and $\delta_4 =$

$\partial DQI/\partial EIQI$. The advantage of using nonparametric technology in this context is that we are able to estimate the first order derivative δ_i for every country in the sample. Thus we are able to gain a deeper understanding of the relationship between development and institutions, for the countries included in our dataset. Out of all the LLS estimates of δ_i , 28% are significant at 1% level, 25% are significant at 5% level and 10% are significant at 10% level of significance. None of the LLS estimates of δ_2 , δ_3 and δ_4 are significant at 10%.

The LLS estimates are insignificant for most countries even at 10% levels. This is not surprising since the analysis till this stage disregards the information contained in the variance covariance matrix Ω . Classical techniques assume that the stochastic disturbances have constant variance. However, as long as $\Omega \neq \sigma^2 I$, the LLS estimates are not efficient. To deal with the problem, we compute the LLGLS estimates. We represent the LLGLS estimates of the first order derivatives by $\gamma = \{\gamma_1 \gamma_2 \gamma_3 \gamma_4\}$, where $\gamma_1 = \partial DQI/\partial IQI$, $\gamma_2 = \partial DQI/\partial PIQI$, $\gamma_3 = \partial DQI/\partial SIQI$ and $\gamma_4 = \partial DQI/\partial EIQI$. All LLGLS estimates of γ_1 are significant at the 1% level. Out of the remaining LLGLS estimates, 85% of estimates of γ_2 , 27% of estimates of γ_3 and 12% of γ_4 are significant at the 1% level. Some nonparametric estimates are negative. We present all LLS estimates in table 1 and all LLGLS estimates in table 2. We observe that about 80% of all LLS and LLGLS estimates indicate a positive relationship between development and institutional indicators.

As explained by Robinson (1988), a reasonable parametric model affords precise inferences, a badly misspecified one, possibly misleading inferences, while a nonparametric model is associated with greater robustness and lesser precision. Therefore we employ the intermediate strategy of conducting a semiparametric regression. We estimate the semiparametric models in (7) to (10). We represent the semiparametric estimates of the first order derivatives by $\phi = \{\phi_1 \phi_2 \phi_3 \phi_4\}$, where $\phi_1 = \partial DQI/\partial IQI$, $\phi_2 = \partial DQI/\partial PIQI$, $\phi_3 = \partial DQI/\partial SIQI$ and $\phi_4 = \partial DQI/\partial EIQI$. Out of the semiparametric estimates, 94% of estimates of

ϕ_1 , 30% of estimates of ϕ_2 , 87% of estimates of ϕ_3 and 32% of estimates of ϕ_4 are significant at the 1% level. 84% of all semiparametric estimates indicate a positive relationship between development and institutional indicators. Details are available in table 3. Figure 1 compares the LLS estimates of $\delta = \{\delta_1 \delta_2 \delta_3 \delta_4\}$, figure 2 compares the LLGLS estimates of $\gamma = \{\gamma_1 \gamma_2 \gamma_3 \gamma_4\}$ and figure 3 compares the semiparametric estimates of $\phi = \{\phi_1 \phi_2 \phi_3 \phi_4\}$.

4: Conclusions

The results of the nonparametric model of our paper support the notion that in general “Institutions Rule”. The level of institutional quality is an important determinant of development quality at the cross-country level for the time period of 25 years. It is possible that countries with better institutional quality are in a better position to reap benefits from trade integration and geography. On the other hand, countries with weak institutional quality find it difficult to enhance their overall development level. Overall, our preliminary results indicate that in addition to significance of institutions, the role of economic policies and geography are also key in determining the level of development. Hence, the level of institutions, economic policies and geography are the three key determinants of the differential levels of development across countries. Their relative significance in explaining development quality depends on the exact stage of development of the country.

In general, we believe that the novelty of our paper is the use of non-parametric technique to understand the importance of institutions in development of a country. Our results strongly support the findings of Basu (2008) in the context of parametric framework. Future research will investigate the development-institution relationship further by estimating a fixed/ random effects nonparametric model. The model can be enhanced by adding more covariates, which can capture the channels followed by countries to climb up in the ladder of development.

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Table 1: Local Linear Least Squares Derivatives

CCODE	δ_1	δ_2	δ_3	δ_4	CCODE	δ_1	δ_2	δ_3	δ_4
AGO	1.078*	0.202	0.058	0.149	KEN	1.659	4.168	0.194	0.258
ALB	1.059**	0.285	0.194	0.328	KOR	0.918*	4.055	0.128	0.094
ARE	0.901	0.25	-1.819	-1.71	KWT	0.931	0.53	1.221	-0.529
ARG	1.67*	3.322	0.235	-0.082	LBR	1.066*	-0.188	-0.168	-0.047
AUS	1.497*	4.134	0.358	-0.025	LKA	1.713**	4.341	0.083	0.044
AUT	0.769*	4.421	1.082	0.231	LUX	0.948	0.107	0.355	0.161
BEL	0.852*	4.493	1.527	0.049	MAR	4.015	1.948	-0.194	-0.096
BGD	1.698	3.633	0.376	-0.062	MDG	3.994	2.455	0.013	-0.027
BGR	0.956	0.536	0.315	-2.342	MEX	2.139*	5.202	-0.285	-0.34
BHR	0.996**	0.068	-0.194	-0.225	MLI	4.142	2.245	0.068	0.038
BOL	1.713**	4.219	-0.188	-0.191	MOZ	1.724	3.912	-0.037	-0.116
BRA	1.672**	3.51	0.176	0.253	MWI	1.743**	4.321	-0.176	-0.18
BWA	1.078*	0.33	0.18	0.269	MYS	1.628**	4.452	0.163	0.118
CAN	1.44*	4.17	0.124	0.083	NER	1.079**	0.173	-0.004	-0.354
CHE	0.967	0.268	0.15	-0.001	NGA	3.14	5.373	-0.184	-0.169
CHL	2.237	4.318	0.055	0.037	NIC	1.726**	4.275	0.071	0.064
CHN	1.043**	0.391	-0.222	-0.474	NLD	1.005*	0.272	0.09	-0.004
CIV	3.66	3.072	-0.191	-0.127	NOR	0.428	4.406	0.159	0.1
CMR	3.346	5.083	-0.08	-0.066	NZL	1.497*	4.211	0.093	0.148
COL	1.706**	4.112	-0.185	-0.156	OMN	1.014**	0.225	0.161	0.152
CRI	1.569*	4.365	0.097	0.071	PAK	3.38	4.133	-0.031	-0.029
DNK	0.987	0.245	-0.128	-0.047	PAN	1.669**	4.428	0.151	0.242
DOM	1.652*	4.311	0.193	0.268	PER	2.95	4.871	-0.043	-0.024
DZA	1.624*	4.083	-0.097	-0.211	PHL	3.66	2.251	0.021	0.005
ECU	1.685*	4.298	-0.158	-0.217	PNG	1.079*	0.405	0.037	0.354
EGY	1.739*	4.141	0.072	0.106	POL	0.949	-0.494	1.666	0.011
ESP	1.035**	0.341	-0.048	-0.01	PRT	1.128*	4.379	1.187	-0.178
ETH	1.688	3.523	0.076	0.113	PRY	1.728*	4.238	0.142	0.231
FIN	0.321	4.372	0.114	0.26	ROM	1.014	0.378	0.277	0.029
FRA	2.883	3.94	0.005	-0.03	SAU	0.982	0.357	0.037	0.104
GAB	1.058**	0.325	0.18	0.113	SDN	1.077*	0.015	0.147	-0.326
GBR	1.862*	5.015	-0.144	-0.252	SEN	4.093	2.156	0.011	-0.027
GHA	3.828	2.814	0.046	0.025	SGP	0.761	3.262	0.01	-0.019
GIN	1.077*	0.343	0.062	0.133	SLV	3.827	2.27	0.077	-0.015
GNB	1.067*	0.4	0.024	0.103	SWE	0.318	4.395	0.113	0.27
GRC	1.041**	0.339	-0.146	-0.945	SYR	1.078*	0.457	0.178	0.356
GTM	1.729*	4.122	-0.185	-0.199	TGO	1.075**	0.436	0.02	0.163
GUY	1.667	4.178	0.176	0.281	THA	1.551*	4.389	-0.018	-0.182
HND	3.869	2.087	0.015	-0.002	TTO	1.609*	4.626	-0.173	-0.161
HTI	4.087	2.449	0.019	-0.029	TUN	1.722*	4.271	0.159	0.245
HUN	1.051**	0.102	0.195	-0.095	TUR	1.067**	0.311	0.075	0.072
IDN	1.748*	4.239	-0.205	-0.255	TZA	1.697**	4.46	-0.213	-0.282
IND	1.676**	3.43	0.088	0.14	UGA	4.055	2.662	-0.037	-0.019
IRL	0.811*	4.483	0.034	0.228	URY	1.672**	4.116	0.077	0.074
IRN	1.066*	0.257	0.297	-0.01	USA	1.04*	4.215	-0.103	-0.129
ISL	0.394	4.013	0.116	0.086	VEN	1.668*	4.249	0.117	0.087
ISR	1.008*	0.22	0.147	0.219	VNM	1.07*	-0.044	0.906	-0.747
ITA	0.851*	4.282	0.448	-0.166	ZAF	3.378	3.275	0.046	0.025
JAM	1.645**	4.372	0.176	0.265	ZAR	1.078*	0.431	0	0.128
JOR	1.067**	0.262	0.203	0.071	ZMB	1.725**	4.338	0.203	0.157
JPN	1.003*	0.265	0.134	0	ZWE	1.662	3.999	-0.699	-0.705

Notes: *Significant at 1 percent, **Significant at 5 percent and *** Significant at 10 percent.

Table 2: Local Linear Generalized Least Squares Derivatives

ccode	γ_1	γ_2	γ_3	γ_4	ccode	γ_1	γ_2	γ_3	γ_4
AGO	1.07*	0.32*	0	0.14*	KEN	1.83*	4.05*	0.27*	0.16*
ALB	1.08*	0.29*	0.19*	0.27*	KOR	0.8*	3.87*	0.1*	-0.02
ARE	0.86*	0.31*	-1.89	-1.7	KWT	0.92*	0.49*	1.25*	-0.35
ARG	1.84*	3.31*	0.06	0.03	LBR	1.07*	-0.31*	-0.17	0.02
AUS	1.74*	4.05*	0.41*	-0.05	LKA	2*	4.46*	0.34**	-0.03
AUT	0.75*	4.61*	0.78	0.41	LUX	0.96*	0.34**	0.43*	0.09
BEL	0.8*	4.59*	1.31*	0.27	MAR	4.39*	2.68	-0.28	-0.14
BGD	1.96*	3.54*	0.33*	0.07	MDG	4.21*	3.56**	0.02	0.01
BGR	0.93*	0.47*	0.23***	-2.77	MEX	2.51*	5.24*	-0.3	-0.32
BHR	1.01*	0.06*	-0.22	-0.11	MLI	4.32*	3.36***	0.06*	0.02
BOL	2*	4.18*	-0.2	-0.11	MOZ	1.87*	3.74*	-0.06	-0.12
BRA	1.88*	3.46*	0.26*	0.15**	MWI	2.14*	4.4*	-0.17	-0.12
BWA	1.07*	0.4*	0.26*	0.12**	MYS	1.75*	4.51*	0.09	0.09
CAN	1.72*	4.09*	0.39*	0	NER	1.07*	0.19*	0.11	-0.35
CHE	0.96*	0.31*	0.25**	0	NGA	3.26*	5.43*	-0.28	-0.2
CHL	1.98*	4.65*	0.07*	0.04*	NIC	2.05*	4.3*	0.3**	-0.05
CHN	1.07*	0.34*	-0.26	-0.36	NLD	1.01*	0.31*	0.32**	0
CIV	4.2*	4.2*	-0.28	-0.16	NOR	0.43*	4.57*	0.11**	0.08
CMR	3.43*	5.02*	-0.11	-0.08	NZL	1.79*	4.16*	0.28**	0
COL	1.98*	4.05*	-0.18	-0.13	OMN	1.02*	0.3*	0.11**	0.11*
CRI	1.73*	4.52*	0.36**	-0.02	PAK	4.08*	4.26*	-0.06	-0.04
DNK	1*	0.32*	-0.07	0	PAN	1.86*	4.47*	0.24*	0.14***
DOM	1.78*	4.4*	0.26*	0.24*	PER	3.64*	4.8*	-0.03	0
DZA	1.85*	3.87*	-0.04	-0.09	PHL	4.1*	3.41***	0.03	0.01
ECU	1.88*	4.34*	-0.13	-0.13	PNG	1.08*	0.34*	-0.01	0.41*
EGY	2.05*	4.02*	0.25***	-0.05	POL	0.93*	-0.78*	1.8*	-0.07
ESP	1.04*	0.32*	-0.1	-0.02	PRT	1.14*	4.51*	1.39*	-0.11
ETH	1.7*	3.4*	0.28**	-0.04	PRY	2.03*	4.21*	0.24**	0.13***
FIN	0.28*	4.46*	0.12*	0.19*	ROM	1.01*	0.33*	0.3*	0.1*
FRA	3.67*	4.24*	0.02	-0.03	SAU	1*	0.31*	-0.01	0.07
GAB	1.06*	0.31*	0.1	0.06	SDN	1.08*	0.15	0.05	-0.25
GBR	2.13*	5.06*	-0.23	-0.28	SEN	4.37*	3.27***	0.02	0.01
GHA	4.07*	3.67*	0.05**	0	SGP	0.56*	3.7*	0.03	0.01
GIN	1.08*	0.32*	-0.04	0.1**	SLV	4.25*	3.41***	0.12*	0.02
GNB	1.08*	0.34*	-0.01	-0.05	SWE	0.3*	4.53*	0.11*	0.15*
GRC	1.04*	0.32*	-0.11	-0.8	SYR	1.07*	0.38*	0.14***	0.35*
GTM	2.06*	4.02*	-0.2	-0.11	TGO	1.08*	0.37*	0	0.09***
GUY	1.84*	4.06*	0.2**	0.29**	THA	1.73*	4.56*	-0.07	-0.11
HND	4.35*	3.17	0.03	0.01	TTO	1.72*	4.78*	-0.23	-0.11
HTI	4.22*	3.67**	0.02	0.01	TUN	2.01*	4.28*	0.26*	0.14**
HUN	1.05*	0.26	0.18*	-0.13	TUR	1.08*	0.32*	0.03	0.05
IDN	2.17*	4.23*	-0.28	-0.15	TZA	1.97*	4.53*	-0.25	-0.22
IND	1.88*	3.35*	0.25***	-0.03	UGA	4.15*	3.7**	-0.06	-0.05
IRL	0.74*	4.62*	0.08	0.13***	URY	1.88*	4*	0.32**	-0.03
IRN	1.06*	0.32*	0.41*	0.01	USA	1.12*	4.2*	-0.12	-0.14
ISL	0.35*	3.89*	0.33**	-0.04	VEN	1.83*	4.24*	0	0.07
ISR	1.01*	0.35*	0.32*	0.16	VNM	1.08*	-0.04*	1.33*	-0.67
ITA	0.75*	4.28*	0.4***	-0.11	ZAF	4.1*	4.06*	0.04**	0.01
JAM	1.76*	4.51*	0.25*	0.15*	ZAR	1.08*	0.36*	0	0.05
JOR	1.08*	0.35*	0.39**	0.04	ZMB	2.05*	4.45*	0.21*	0.08
JPN	1.01*	0.32*	0.02	0	ZWE	1.82*	3.82*	-0.66	-0.7

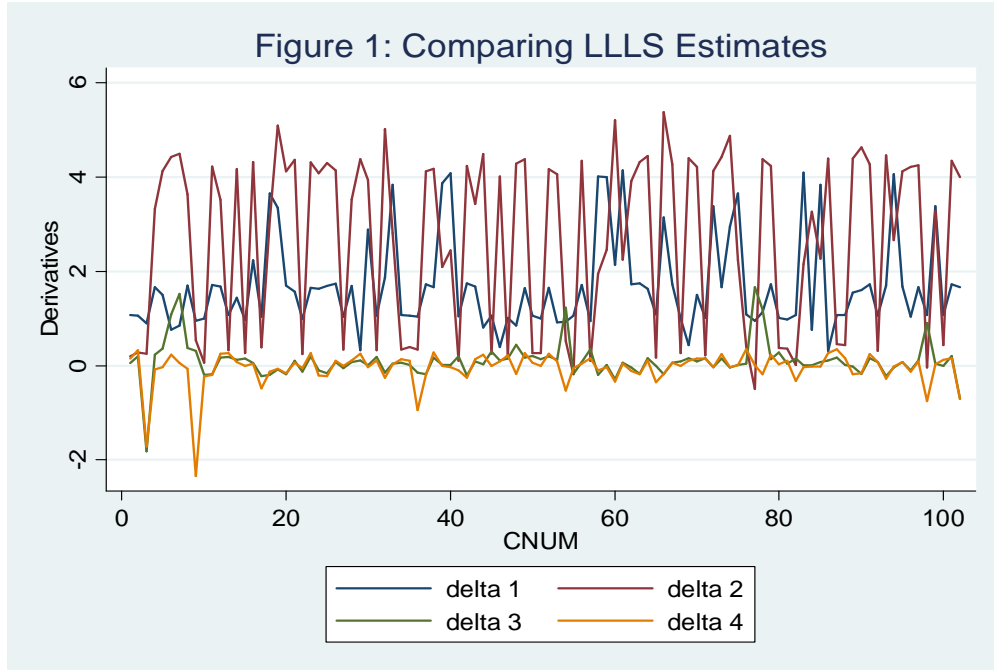
Notes: *Significant at 1 percent, **Significant at 5 percent and *** Significant at 10 percent.

Table 3: Derivatives from the Semiparametric Model

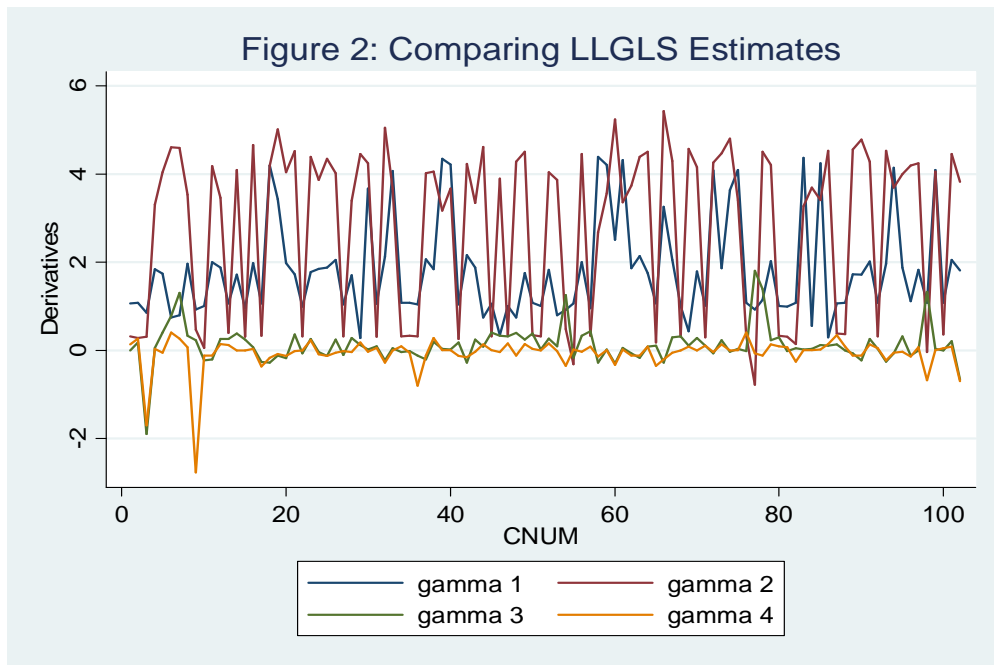
<i>CCODE</i>	ϕ_1	ϕ_2	ϕ_3	ϕ_4	<i>CCODE</i>	ϕ_1	ϕ_2	ϕ_3	ϕ_4
ALB	0.4874	0.45*	0.28*	-0.00023	KOR	0.3462	0.14	0.60**	-0.00006
ARE	0.0382	-0.08	-0.07	0.00147***	KWT	0.2665	-0.08	0.27*	0.00008
ARG	1.3999	0.19**	0.65*	0.05216*	LBR	0.3922	-0.45	-0.05	0.00092
AUS	0.6566	0.21*	0.67*	0.00396*	LKA	1.6356	0.14	0.60*	0.00188
AUT	0.3588	0.16***	0.58***	0.00035	LUX	0.1896	0.12	0.68*	-0.00864
BEL	0.3563	0.17**	0.77*	-0.00033	MAR	1.5302	0.18**	0.32*	0.00115
BGD	1.6748	0.1	0.43*	0.06714*	MDG	1.6522	0.16***	0.32*	0.01746*
BGR	0.4798	-0.11	0.28*	-0.00295	MEX	0.9184	0.18**	0.33*	-0.00749
BOL	1.6169	0.17***	0.48*	0.06387*	MLI	1.6676	0.08	0.29*	0.05277*
BRA	1.2871	0.18**	0.66*	-0.00032	MOZ	1.7502	-0.05	0.62*	0.00494*
BWA	0.4044	0.22*	0.26*	0.00372*	MWI	1.663	-0.01	0.64*	-0.00026
CAN	0.6272	0.19**	0.67*	0.00026	MYS	1.4944	0.13	0.68*	0.00033
CHE	0.1118	-0.09	0.68*	0.00309*	NER	0.3569	0.21*	0.11**	0.00007
CHL	0.3255	0.21*	0.37*	0.00049	NGA	1.794	0.1	0.33*	0.00738*
CHN	0.2672	0.03*	0.68*	0.00054	NIC	1.6605	0.15***	0.58*	-0.00032
CIV	1.6641	0.16***	0.33*	0.00238***	NLD	0.4163	0.24	0.68*	-0.0058
CMR	1.7837	0.15***	0.33*	-0.00033	NOR	0.4359	0.22*	-0.05	-0.00025
COL	1.556	0.15***	0.58*	-0.00027	NZL	0.6378	0.19**	0.67*	0.00347**
CRI	1.0229	0.22*	0.68*	-0.0003	OMN	0.4748	0.04	0.35*	0.00386*
DNK	0.2979	0.28	0.68*	-0.00386	PAK	1.6633	0.12	0.32*	0.00124
DOM	1.3097	0.20*	0.65*	0.06988*	PAN	1.6467	0.19**	0.54*	0.00066
DZA	1.2481	-0.01	0.60*	0.00391*	PER	1.1309	0.12	0.32*	0.00503*
ECU	1.4293	0.18**	0.67*	0.01203*	PHL	1.0441	0.15***	0.33*	0.00029
EGY	1.6805	0.09	0.59*	0.00015	PNG	0.3921	0.14*	-0.15*	-0.00025
ESP	0.4639	0.65*	0.68*	-0.00862	POL	0.4679	-0.1	0.23*	-0.00023
ETH	1.7157	-0.04	0.59*	0.00745*	PRT	0.2659	0.06	0.41*	-0.00028
FIN	0.4578	0.21*	0.2	0.00577*	PRY	1.6636	0.15***	0.53*	0.02543*
FRA	0.9036	0.21*	0.64*	0.00045	ROM	0.4391	-0.05	0.15*	0.00552*
GAB	0.4313	-0.09	0.09*	-0.00007	SAU	0.3428	0.04	-0.14**	-0.00019
GBR	0.9044	0.20*	0.67*	0.00045	SDN	0.3533	0.1	0.18*	0.00115
GHA	1.6629	0.13	0.33*	0.01004*	SEN	1.6603	0.14	0.32*	0.00839*
GIN	0.3543	0.04*	0.45*	0.00644*	SGP	0.4289	0.22*	0.55*	-0.00762
GNB	0.3368	0.1	-0.12**	0.00033	SLV	1.2358	0.1	0.33*	0.00016
GRC	0.4466	0.58*	0.68*	0.00402*	SWE	0.463	0.22*	0.21	-0.00005
GTM	1.6635	0.14***	0.48*	0.05646*	SYR	0.3909	-0.12	0.53*	0.00049
GUY	1.6428	-0.08	0.04	-0.00026	TGO	0.3513	0.02	0.09**	0.00028
HND	1.2397	0.11	0.32*	0.00528*	THA	0.8912	-0.01	0.30*	-0.00028
HTI	1.6726	0.04	0.30*	-0.00033	TTO	0.9976	0.21*	0.47*	0.00081
HUN	0.4295	-0.09	0.18*	-0.00013	TUN	1.6641	0.14	0.62*	0.00028
IDN	1.7109	0.09	0.51*	0.03679*	TUR	0.4635	0.17**	0.66*	0.00298*
IND	1.4126	0.15***	0.66*	0.00006	TZA	1.6637	-0.02	0.67*	0.00045
IRL	0.3592	0.13	0.74*	-0.00863	UGA	1.7402	0.11	0.29*	0.00357**
IRN	0.4555	-0.08	0.89*	0.00029	URY	1.3731	0.20*	0.61*	-0.0001
ISL	0.4412	-0.11	0.25*	-0.0001	USA	0.3539	0.20*	0.67*	-0.00632
ISR	0.4736	0.58*	0.68*	-0.00029	VEN	1.2954	0.19**	0.66*	0.04696*
ITA	0.3569	0.15***	0.76*	0.00009	VNM	0.3425	0.07	0.28*	0.00067
JAM	1.3915	0.21*	0.60*	0.02726*	ZAF	1.4008	0.20*	0.35*	0.00124
JOR	0.4598	-0.03	0.67*	0.00577*	ZAR	0.3554	-0.08	-0.72	0.00434*

JPN	0.4103	-0.29	0.68*	0.0003	ZMB	1.654	-0.01	0.68*	-0.00032
KEN	1.5811	0.04*	0.68*	0.00598*	ZWE	1.4227	-0.03	0.66*	-0.00012

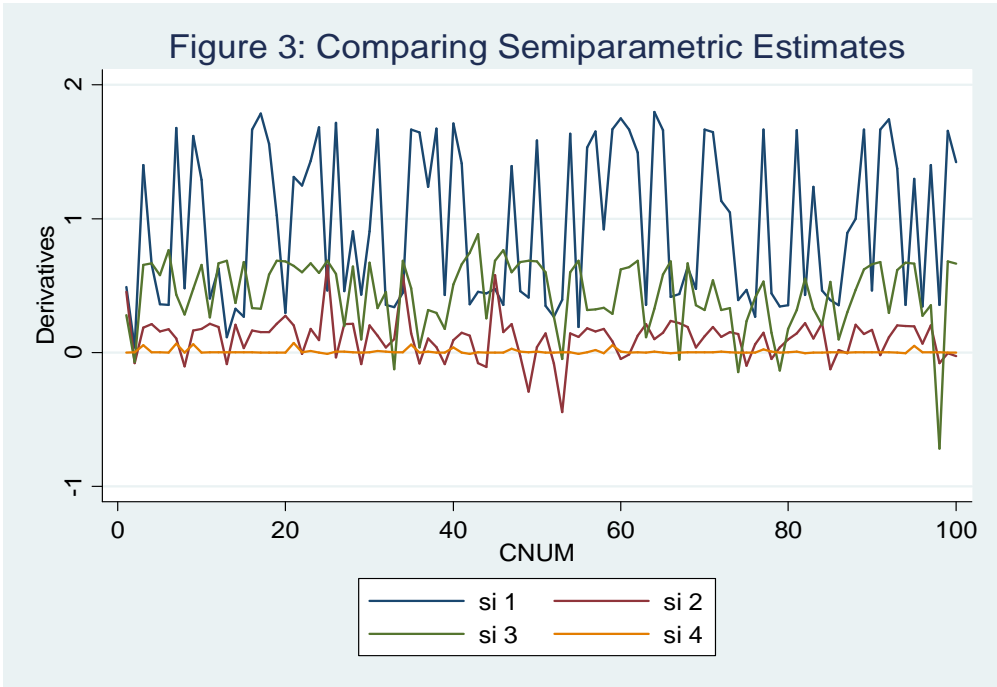
Notes: *Significant at 1 percent, **Significant at 5 percent and *** Significant at 10 percent.



* Note: delta 1 = δ_1 , delta 2 = δ_2 , delta 3 = δ_3 and delta 4 = δ_4 . See A3 for CNUM.



*Note: gamma 1 = γ_1 , gamma 2 = γ_2 , gamma 3 = γ_3 and gamma 4 = γ_4 . See A3 for CNUM.



*Note: si 1 = ϕ_1 , si 2 = ϕ_2 , si 3 = ϕ_3 and si 4 = ϕ_4 . See A3 for CNUM.

Annex Tables

Table A1: List of countries in sample

Country Code	OECD (22)	Country Code	Latin America (22)
AUS	Australia	BOL	Bolivia
JPN	Japan	COL	Colombia
NZL	New Zealand	CRI	Costa Rica
GRC	Greece	DOM	Dominican Republic
PRT	Portugal	ECU	Ecuador
CAN	Canada	GTM	Guatemala
USA	United States	GUY	Guyana
AUT	Austria	JAM	Jamaica
BEL	Belgium	PER	Peru
CHE	Switzerland	PRY	Paraguay
DNK	Denmark	SLV	El Salvador
ESP	Spain	HND	Honduras
FIN	Finland	HTI	Haiti
FRA	France	NIC	Nicaragua
GBR	United Kingdom	ARG	Argentina
IRL	Ireland	BRA	Brazil
ISL	Iceland	CHL	Chile
ITA	Italy	MEX	Mexico
LUX	Luxembourg	PAN	Panama
NLD	Netherlands	TTO	Trinidad and Tobago
NOR	Norway	URY	Uruguay
SWE	Sweden	VEN	Venezuela
Country Code	Sub-Saharan Africa (26)	Country Code	Asia and Pacific (13)
AGO	Angola	BGD	Bangladesh
BWA	Botswana	CHN	China
CIV	Cote d'Ivoire	IDN	Indonesia
CMR	Cameroon	IND	India
ETH	Ethiopia	KOR	Korea, Rep.
GAB	Gabon	LKA	Sri Lanka
GHA	Ghana	MYS	Malaysia
GIN	Guinea	PAK	Pakistan
GNB	Guinea-Bissau	SGP	Philippines
KEN	Kenya	SGP	Singapore
LBR	Liberia	THA	Thailand
MDG	Madagascar	VNM	Vietnam
MLI	Mali	PNG	Papua New Guinea
MOZ	Mozambique		
MWI	Malawi	Country Code	Middle East and North Africa (13)
NER	Niger	ARE	United Arab Emirates
NGA	Nigeria	ISR	Israel
SDN	Sudan	KWT	Kuwait
SEN	Senegal	IRN	Iran, Islamic Rep.
TGO	Togo	JOR	Jordan
TZA	Tanzania	SYR	Syrian Arab Republic
UGA	Uganda	BHR	Bahrain
ZAF	South Africa	OMN	Oman
ZAR	Congo, Dem. Rep.	SAU	Saudi Arabia
ZMB	Zambia	DZA	Algeria
ZWE	Zimbabwe	EGY	Egypt, Arab Rep.
		MAR	Morocco
		TUN	Tunisia
Country Code	EU and Other Europe (6)		
ALB	Albania		
BGR	Bulgaria		
ROM	Romania		
HUN	Hungary		
POL	Poland		
TUR	Turkey		

Source: United Nations and World Bank

Table A2: Development Quality Index (DQI) and Institutional Quality Index (IQI): Definition and Sources of Indicators

Economic DQI	Economic IQI
GDP per capita (PPP, \$ international 2000)	Legal and property rights ³
Telephone mainlines (per 1,000 people)	Law and Order ^{1a}
Television sets (per 1,000 people)	Bureaucratic Quality ^{1a}
Radios (per 1,000 people)	Corruption ^{1a}
Electric power consumption (kwh per capita)	Democratic Accountability ^{1a}
Energy use (kg of oil equivalent per capita)	Government Stability ^{1a}
	Independent Judiciary ²
	Regulation ³
Health DQI	Social IQI
Life expectancy at birth, total (years)	Press Freedom ³
Mortality rate, infant (per 1,000 live births)	Civil Liberties ³
Physicians (per 1,000 people)	Physical Integrity Index ⁴
Immunization, DPT (% of children ages 12-23 months)	Empowerment Right Index ⁴
CO2 emissions (metric tons per capita)	Freedom of Association ⁴
	Women's Political Rights ⁴
	Women's Economic Right ⁴
	Women's Social Right ⁴
Knowledge DQI	Political IQI
Literacy rate, adult total (% of people ages 15 and above)	Executive Constraint ⁶
School enrolment, primary (% gross)	Political Rights ³
School enrolment, secondary (% gross)	Index of Democracy ⁵
Total number of years in schools ¹	Polity Score ⁶
	Lower Legislative ²
	Upper Legislative ²
	Independent Sub-federal Units ²

Note. For DQI, data obtained from the World Development indicators CD-ROM 2006, World Bank; and ¹Barro and Lee 2000 dataset, ^{1a}PRS Group (2005) ICRG database; ²POLCON Henisz Dataset; ³Economic Freedom Index dataset, Freedom House; ⁴CIRI Human Rights Data Project; ⁵PRIO Dataset; ⁶Polity IV Project

Table A3: Countries and CNUM.

CCODE	CNUM	CCODE	CNUM	CCODE	CNUM
AGO	1	GNB	35	NOR	69
ALB	2	GRC	36	NZL	70
ARE	3	GTM	37	OMN	71
ARG	4	GUY	38	PAK	72
AUS	5	HND	39	PAN	73
AUT	6	HTI	40	PER	74
BEL	7	HUN	41	PHL	75
BGD	8	IDN	42	PNG	76
BGR	9	IND	43	POL	77
BHR	10	IRL	44	PRT	78
BOL	11	IRN	45	PRY	79
BRA	12	ISL	46	ROM	80
BWA	13	ISR	47	SAU	81
CAN	14	ITA	48	SDN	82
CHE	15	JAM	49	SEN	83
CHL	16	JOR	50	SGP	84
CHN	17	JPN	51	SLV	85
CIV	18	KEN	52	SWE	86
CMR	19	KOR	53	SYR	87
COL	20	KWT	54	TGO	88
CRI	21	LBR	55	THA	89
DNK	22	LKA	56	TTO	90
DOM	23	LUX	57	TUN	91
DZA	24	MAR	58	TUR	92
ECU	25	MDG	59	TZA	93
EGY	26	MEX	60	UGA	94
ESP	27	MLI	61	URY	95
ETH	28	MOZ	62	USA	96
FIN	29	MWI	63	VEN	97
FRA	30	MYS	64	VNM	98
GAB	31	NER	65	ZAF	99
GBR	32	NGA	66	ZAR	100
GHA	33	NIC	67	ZMB	101
GIN	34	NLD	68	ZWE	102